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A CROSS-COUNTRY ESTIMATION OF THE ELASTICITY OF
SUBSTITUTION BETWEEN LABOR AND CAPITAL IN
MANUFACTURING INDUSTRIES*

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ABSTRACT

This paper presents a simple methodology to estimate the elasticity of substitution between labor and capital for firms operating in perfectly competitive factor markets with constant-elasticity-of-substitution technologies. It is applied to a cross-country sample of 28 3-digit ISIC manufacturing industries. The econometric procedure relies on measures of sectorial capital stock, that are estimated for 34 countries in 1990. Unlike previous studies, the estimates are compatible with international technology differences. The results reveal that in most industries the elasticity of substitution is close to one. However, the null hypothesis of Cobb-Douglas production functions is in general rejected.

* I acknowledge the comments by two anonymous referees. All remaining errors are my responsibility.

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RESUMEN

El trabajo presenta una metodología simple para estimar la elasticidad de sustitución entre trabajo y capital de firmas que operan en un mercado de factores perfectamente competitivos con tecnologías con elasticidad de sustitución constante. La metodología se aplica a una muestra de países a través de 28 industrias manufactureras (3 dígitos ISIC). La estimación requiere medidas sectoriales de stock de capital, que es estimado para 34 países en 1990. Contrariamente a estudios anteriores, la estimación es compatible con diferencias tecnológicas entre países. Los resultados muestran que en la mayoría de las industrias la elasticidad de sustitución es cercana a 1. Sin embargo, la hipótesis nula de $\sigma_{LK} = 1$ es rechazada en la mayoría de los casos.

1. INTRODUCTION

This paper presents a cross-country estimation of the elasticity of substitution between labor and capital – σ_{LK} for 28 3-digit ISIC manufacturing industries. The econometric estimation, performed for 34 countries in 1990, is based on the predictions of the relationship between relative factor prices and relative factor intensities that follow from the optimizing behavior of firms operating in perfectly-competitive factor markets with technologies that can be represented with CES production functions. It is assumed that there exists many identical firms within each country-industry pair, so that each firm takes factor prices as given and chooses its capital-labor ratio according to traditional first-order conditions. An estimation of the elasticity of substitution between labor and capital is relevant to study the effects on factor demand, wages, employment and income distribution of factor accumulation, technological change and international trade.

The main contribution of the paper is that the industry-specific estimates of σ_{LK} are compatible with international hicks-neutral technology differences. As discussed in the paper, alternative approaches rely on very restrictive assumptions of technology differences across countries. Traditional studies are based on either direct estimates of production functions or first-order conditions of firms' maximization problems. In the former case, estimations of σ_{LK} that allow for cross-country technology differences are limited by the degrees of freedom available. In the latter case, scarce data on sectorial capital stock across countries have lead researchers to develop methodologies that impose strong restrictions on cross-country technology differences. The methodology proposed in this paper assumes optimizing behavior by firms. It therefore belong to the second group. However, the availability of detailed data on capital stock allows us to produce estimates of the elasticity of substitution with less restrictive assumptions on international technology differences. More specifically, the results are consistent with hicks-neutral international technology differences. This is an important point, as many

researchers have emphasized the role of international technology differences in explaining international factor price differences (see Treffer (1993)).

Also, the results are consistent with cross-country differences in returns to scale. This is because the first order conditions for firms operating in perfectly competitive factor markets and CES production functions are the same regardless of their returns to scale under the assumption that the elasticity of substitution between labor and capital is the same across countries in any given industry. The presence of imperfect competition in response to increasing returns internal to the firms does not imply that the firm hires factors until the ratio of their prices is equal to the ratio of their marginal productivity.

The estimation for only two production factors is limited by the availability of data. This is of course a disputable approach, as we may think that factors of production in manufacturing industries could be classified in broader categories. The interpretation of such elasticity is subject then to the feasibility of aggregating different types of labor into one category. (See Berndt and Christensen (1973) for a discussion on the conditions for aggregating factors.) Although a priori one may think that aggregating skilled labor and capital is more viable than aggregating different employment categories, evidence presented for Chile by Corbo and Meller (1982) suggests that this is not necessarily the case. They argue that for most manufacturing industries there exists some evidence that building an aggregate of unskilled and skilled labor is more reasonable than aggregating skilled labor and capital. These results are based on estimates of translog production functions for 4-digit manufacturing industries. I bring this conclusion to the paper and consider that the aggregation of labor into one broad category is reasonable.

An additional criticism to this approach is that it does not control for differences in the levels of efficiency of factors across countries. However, it is well known that productivity differences can be adequately represented through differences in the efficiency units of factors or by differences in the production functions themselves. Treffer (1993) represents productivity differences in the former way. I do it in the latter way. Nevertheless, they are equivalent in the sense that for any distribution of differences in production functions, there is a vector of differences in efficiency units of factors that yield similar international factor prices.

The paper is divided as follows. Section 2 presents the methodology, and a comparison with alternative approaches to the estimation of σ_{LK} . Section 3 discusses briefly the data. Section 4 reports the results and compares them with other results in the literature. Section 5 concludes.

2. METHODOLOGY

Consider a production function with constant elasticity of substitution between two factors of production labor L and capital K

$$(1) \quad X_{ic} = f(L_{ic}, K_{ic}) = A_{ic} (a_{ic} L_{ic}^{\rho_{ic}} + b_{ic} K_{ic}^{\rho_{ic}})^{1/\tau_{ic}}$$

X_{ic} refers to real value-added of good i in country c ; L_{ic} and K_{ic} represent labor and capital inputs in sector i and $\rho_i \leq 1$. The technology parameters are given by A_{ic} , a_{ic} , b_{ic} , ρ_i and τ_{ic} . Note that we are allowing for a and b to differ across countries. As discussed below, this feature will make our estimation of a common σ_i across countries relevant even in the presence of cross-country Hicks-neutral technology differences. Notice that the technology exhibits constant-returns-to-scale if $\tau_{ic} = \rho_i$. If $\tau_{ic}/\rho_i > 1$, we are in the presence of increasing returns internal to the firm.

The first order conditions of the maximization process of each firm that take factor prices as given are

$$(2) \quad w_{ic} = A_{ic} (a_{ic} L_{ic}^{\rho_i} + b_{ic} K_{ic}^{\rho_i})^{1-\tau_{ic}/\tau_{ic}} a_{ic} L_{ic}^{\rho_i-1} / \tau_{ic}$$

$$(3) \quad r_{ic} = A_{ic} (a_{ic} L_{ic}^{\rho_i} + b_{ic} K_{ic}^{\rho_i})^{1-\tau_{ic}/\tau_{ic}} b_{ic} K_{ic}^{\rho_i-1} / \tau_{ic}$$

where w_{ic} and r_{ic} are nominal factor prices deflated by sectorial prices.¹

As expected, for $\tau_{ic} \neq \rho_i$, the marginal productivity of labor and capital depend on factor proportions as well as the level of factor usage. Combining (2) and (3) we get for each industry i in country c the following relationship between factor prices - w/r and factor usage - K/L .²

$$(4) \quad \ln(w_{ic}/r_{ic}) = \ln(a_{ic}/b_{ic}) + (\rho_i - 1) \ln(L_{ic}/K_{ic})$$

The elasticity of substitution between labor and capital is given by $\sigma_i = 1/(1-\rho_i)$. It measures the ease of substitution along the production isoquant, reflecting the optimizing change in relative factor usage in response to changes in relative factor prices. An increase in the relative cost of labor makes the firm use relatively more capital than before, at any scale of production.

Equation (4) provides the basis for the empirical estimation developed in section 4. I estimate this relationship for each 3-digit ISIC manufacturing industry in a cross-section sample of 34 developed and developing countries in 1990. The absence of the subindex c in the coefficient for $\ln(L_{ic}/K_{ic})$ reveals that the elasticity of substitution in each industry does not vary across countries. However, a common σ_i is compatible with international technology differences. These are

¹ These conditions hold even in the case when product prices differ from marginal costs of production in the presence of imperfect competition.

² Strictly speaking, equation (4) suggests a relationship between factor intensities and relative factor prices at the sectorial level. Without a theory for cross-industry differences in factor prices, (4) reveals that for a common vector of relative factor prices, different industries choose different production techniques. In the empirical section we account for possible differences in factor prices at the sectorial level.

imbedded in differences in a and b . An estimation based on a single intercept for each industry implies that a/b is similar across countries, allowing for hicks-neutral technology differences. This can be seen by dividing (2) and (3) to get

$$(5) \quad \frac{f_L}{f_K} = \frac{w}{r} = \frac{a}{b} \left(\frac{L}{K} \right)^{\rho-1}$$

Differences in a and b that keep the ratio constant imply no change in the optimal K/L for any given relative factor prices. This is exactly what hicks-neutral technology differences imply. With a time-series dimension of the data, it is possible to estimate (4) with country-specific intercepts. However, we only have cross-country data at the industry level. Therefore, the estimation for each sector is performed with a common intercept. This is only consistent with hicks-neutral technological differences.

Also, equation (4) for each industry is valid in countries that have similar ρ_i but differ in τ_{ic} . This implies that some type of differences in returns to scale do not affect expression (4). In particular, we only require that differences in the returns to scale are reflected in differences in τ_{ic} but not ρ_i . This is satisfied as long as the elasticity of substitution is the same regardless of the returns to scale.

A different methodology widely used in the literature is developed by Behrman (1982) and others.³ Like the strategy proposed above, it relies on the optimizing condition of firms operating in perfectly competitive markets. Starting from a CES production function like the one in equation (1) and assuming constant-returns-to scale we can derive the following expression for the first-order condition of maximizing firms⁴

$$(6) \quad \ln(X_{ic}/L_{ic}) = \alpha_{ic} + \sigma_i \ln w_{ic}$$

with $\alpha_{ic} = -\sigma_i (\ln a_{ic} + \ln A_{ic})$. This approach has been used to estimate aggregate measures of elasticity of substitution between labor and capital as well as industry-specific estimations (see Hamermesh (1993)). The main advantage of (6) compared to (4) is that it does not require data on capital stock or return to capital. As (4), it provides a direct estimate of σ . A complete estimation of (6) requires a panel structure of the data in order to allow for different intercepts. Otherwise, very arbitrary assumptions on international technology differences are imposed. Behrman estimates a pooled regression for several countries and 27 3-digit ISIC manufacturing industries using average values for value-added, employment and wages between 1967 and 1973 for more than 70 countries. He includes industry or country dummies to differentiate the elasticity of substitution across different units of analysis. However, due to data limitations (absence of

³ See Arrow, Chenery, Minhas and Solow (1961) and Hamermesh (1993).

⁴ Taking logarithm from the first order condition with respect to labor (equation (2)) we get that $\ln w = \ln a + \ln A + (1-\sigma) \ln(X/L)$ that can be written as (6).

time dimension) he is not able to include a different intercept for each industry-country pair, as (6) suggests. This is a fundamental deficiency of this strategy, as the results are consistent with very restrictive forms of international technology differences.

An alternative approach is the estimation of the production function. The elasticity of substitution can be recovered from the estimates of the technology parameters. In this case, an econometric estimation of (1) allows for less restrictive assumptions regarding the elasticity of substitution. The most flexible alternative is the estimation of translog production functions, that allow for variable elasticity of substitution and not optimizing behavior by firms, avoiding possible specification errors in the presence of noncompetitive elements. Also, econometric estimation of the production function reveals other important parameters, like output elasticities. Berndt and Christensen (1973) discuss the properties of translogarithmic production functions. (See also Griliches and Ringstad (1971).) However, estimations have been developed using cross-section data on firms, establishments or industries where the assumption of common technologies is reasonable. Otherwise, time variation is required to identify differences in the production functions. In the case of this paper, the lack of time-series data would imply an estimation of production functions in each sector under the assumption of no technology differences. In my opinion, this is very restrictive. Therefore, I do not pursue this strategy.⁵ An alternative method for computing σ_{LK} indirectly is to estimate labor demand elasticities (see Hamermesh (1993)). Under optimizing assumptions, it is possible to back-up values for the elasticity of substitution. However, the same pitfalls of the last approach are present. The data requirements or the assumptions needed to estimate labor demands in the presence of technology differences are such that using cross-country data it is very difficult to get sensible estimates.

A final caveat is important when strategies aimed to directly estimate σ are used. They fail to account for the simultaneity of supply and demand. In other words, the relationship from factor prices to factor usage may also go in the opposite direction. This is specially relevant when aggregate estimations are performed, as the supply of factors may depend positively on factor prices. When micro data is used (as in this study), this is less problematic as we can assume that firms within industries face completely elastic factor supplies.

3. DATA

The data is obtained from UNIDO Statistical Database for 180 countries between 1963 and 1996, containing series on employment, value-added, wages and salaries, output and gross fixed capital formation for 28 3-digit industry. The

⁵ Lovell (1973) presents an study on the elasticity of substitution based on the Census of Manufactures across U.S. states in 1958. He concludes that there is little evidence of variable elasticity of substitution (σ as function of factor usage). In most industries, a CES function is more appropriate than VES technologies.

series of real capital stock can be estimated using the capital formation series, an estimation of depreciation rates and adequate investment deflators.

Table 1 reports the measures of capital stock for 1990 calculated for different countries for each 3-digit ISIC manufacturing category, in millions of 1990 US dollars. The series of capital stock was constructed using the yearly series of gross fixed capital formation from 1971 until 1990 in current US dollars. I considered a 5% depreciation rate and the investment deflator series for the United States as the relevant deflator for the capital accumulation series across countries.⁶

Due to data restrictions, I perform the empirical analysis in 1990 because it is year for which the set of countries with capital stock across industries is maximized (34 countries). Tables 2, 3 and 4 report data on employment, nominal value-added and nominal payment of wages and salaries in 1990. These data allow us to calculate measures of L/K for each country/industry pair and estimations of wage-rental rate ratios. The wage rate considered is the average yearly wage of workers (Table 3/ Table 2) and the rental rate is computed as value-added minus labor payments divided by the capital stock (Table 4 - Table 3/Table 1). This method of estimating w and r is similar to that used by Lovell (1973).

4. RESULTS

Table 5 reports the results of regressions of equation (4) for each industry. $(\rho_i - 1)$ represents the coefficient on $\ln(L/K)$, n represents the number of countries included in the regression and $\sigma_i (=1/1-\rho_i)$ is the implicit value for the elasticity of substitution between labor and capital. In all but five industries (Beverages, Tobacco, Petroleum Refineries, Iron & Steel, and non-electrical Machinery) the correlation coefficient between relative factor prices and factor intensity is greater than .7. Figure 1 plots for the 28 ISIC manufacturing industries the cross-country values of $\ln(w_{ic}/r_{ic})$ and $\ln(K_{ic}/L_{ic})$. These graphs show that the high correlations are not driven by outliers, revealing genuine economic relationships. Moreover, in most cases the assumption of a common intercept appears to be adequate. This suggests that if technology differences exist, a hicks-neutral approximation is reasonable. The null hypothesis of Cobb-Douglas technologies ($\rho - 1 = -1$) is rejected in most cases as the last column of Table 5 suggests.⁷

The last two columns of Table 5 report the results of similar regression but using average factor prices (w_c/r_c). The column se^* reports the standard error of the coefficient $1 - \rho$ while σ^* is the estimation of the elasticity of substitution. The correlation coefficient of both series is .39, significant at 5%. Excluding the estimates for Pottery industries (361) and Professional and Scientific Equipment (385), the correlation rises to .67 significant at 1%. These results are not subject to the criticism of endogeneity of relative factor prices.

⁶ The results are not affected by different depreciation rates.

⁷ The p-value reports the probability that $\rho = 0$ that implies $\sigma = 1$. Therefore, low p-values imply the rejection of the Cobb-Douglas hypothesis.

TABLE 1
CAPITAL STOCK FOR 1990
(Millions of 1990 US dollars)

UNIDO Code	Country	Food Products (311/312)	Beverage (313)	Tobacco (314)	Textile (321)	Apparel (322)	Leather (323)	Footwear (324)	Wood (331)	Furniture (332)	Paper Products (341)	Printing and Publishing (342)	Chemical s(352)	Other s(353)	Petroleum Products (354)	Misc. Products (355)	Rubber (356)	Plastic (357)	Pottery (361)	Glass (362)	Other Non- Metallic Products (369)	Iron & Steel (371)	Non- Ferrous Metals (372)	Fabricated Metal Products (381)	Machinery except Electrical (382)	Electrical Machinery (383)	Transport Equipment (384)	Professional Equipment (385)	Other Manufacturing Industries (390)		
2969.0	Austria	..	179.5	1607.5	82.1	212.8	1157.1	..	3498.5	1094.2	208.1	1514.8	1116.1	..	397.1	795.3	98.6	764.1	2068.1	2790.5	645.7	2347.3	2516.6	3610.5	..	234.6	..	202.8	
52.3	Barbados
56.0	Belarus
1290.3	Belgium	1658.2	232.4	3445.0	26.9	438.6	..	1750.3	2512.6	413.6	..	2869.6	3843.2	983.8	18.5
1290.3	Belgium	1658.2	232.4	3445.0	26.9	438.6	..	1750.3	2512.6	413.6	..	2869.6	3843.2	983.8	18.5
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Source: UNIDO Database, FRED and author's calculations.

TABLE 2
EMPLOYMENT LEVEL FOR 1990
(Thousands of workers)

UNIDO Code	Country	Food Products (311/312)	Beverage (313)	Tobacco (314)	Textile (321)	Apparel (322)	Leather (323)	Footwear (324)	Wood (331)	Furniture (332)	Paper Products (341)	Printing Products (342)	Chemicals (351)	Other Chemicals (352)	Rubber Products (353)	Misc. Products (354)	Plastic (355)	Pottery (356)	Glass (361)	Other Metallic Products (369)	Iron & Steel (371)	Non- Ferrous Metals (372)	Fabricated Products (381)	Machinery except Electrical (382)	Electrical Machinery (383)	Transport Equipment (384)	Professional & Scientific Equipment (385)	Other Manufacturing Industries (390)		
40	Austria	52.8	12.7	1.4	37.1	24.1	2.8	8.4	89.3	308	20.0	24.6	20.3	18.7	3.7	1.1	7.1	13.7	3.0	9.0	25.9	36.5	9.6	57.9	730	83.0	31.7	6.8	7.1	
52	Barbados	1.6	0.6	-	0.1	1.1	-	-	0.3	12.8	0.6	0.6	0.4	0.2	-	-	4.6	18.5	3.2	13.4	20.3	39.9	12.0	259.5	0.7	0.1	-	0.1	0.1	
56	Belgium	75.8	12.8	4.9	54.2	36.0	1.0	17.6	12.8	204	17.9	35.0	79.0	31.6	-	-	14.1	18.6	1.8	13.8	40.8	51.0	44.1	13.2	201.2	83.1	-	-	9.5	
100	Bolivia	197.8	24.0	5.3	72.0	94.0	5.1	11.0	102.9	640	114.0	142.0	33.0	68.0	15.0	3.0	23.0	62.0	1.0	12.0	50.0	50.0	44.1	13.0	148.0	123.0	203.9	17.3	19.8	
120	Bulgaria	12.0	1.2	0.2	72.8	68.0	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	11.8	
170	Colombia	79.3	23.3	2.0	52.8	46.8	8.1	15.2	6.3	8.7	11.8	21.9	16.7	26.1	5.4	1.0	6.4	18.6	5.3	7.0	20.5	9.3	2.2	28.5	158	18.2	19.0	3.9	8.7	
196	Cyprus	5.9	1.9	0.4	2.2	11.3	0.9	2.1	1.9	2.3	0.8	1.8	0.1	1.3	0.1	-	0.1	1.2	0.1	0.1	2.4	168.0	24.0	164.0	500.0	189.0	204.0	20.0	46.0	
200	Czechoslovakia	168.0	34.0	4.0	20.0	93.0	25.0	67.0	56.0	590	45.0	28.0	89.0	27.0	23.0	22.0	27.0	8.0	10.0	73.0	75.0	106.0	24.0	164.0	500.0	189.0	204.0	20.0	46.0	
208	Denmark	27.4	7.0	0.8	16.3	11.6	0.0	1.7	0.0	22.2	11.0	55.3	10.6	17.6	0.6	2.8	2.9	12.7	1.9	2.7	4.5	4.4	1.9	45.5	80.7	31.5	25.9	15.3	11.2	
218	Ecuador	27.7	7.4	0.8	16.3	11.6	0.0	1.7	0.0	22.2	11.0	55.3	10.6	17.6	0.6	2.8	2.9	12.7	1.9	2.7	4.5	4.4	1.9	45.5	80.7	31.5	25.9	15.3	11.2	
242	Finland	46.6	5.5	1.1	10.9	14.7	1.4	3.2	30.4	10.9	44.4	37.8	10.3	10.3	3.1	0.7	2.6	7.6	1.3	3.2	15.6	12.8	4.3	31.9	52.4	28.5	27.7	5.8	3.6	
246	Finland	46.6	5.5	1.1	10.9	14.7	1.4	3.2	30.4	10.9	44.4	37.8	10.3	10.3	3.1	0.7	2.6	7.6	1.3	3.2	15.6	12.8	4.3	31.9	52.4	28.5	27.7	5.8	3.6	
250	France	462.5	48.8	5.1	209.9	144.7	19.9	52.5	95.2	835	107.4	233.6	118.3	184.3	23.4	-	94.0	121.2	36.9	70.4	18.9	188.4	48.3	354.4	4398	474.0	535.6	67.9	100.2	
280	Germany, West	576.3	86.5	15.5	228.8	143.2	20.9	50.5	103.9	1417	163.3	180.4	308.0	286.4	23.0	-	98.4	286.2	36.9	70.4	18.9	188.4	48.3	354.4	4398	474.0	535.6	67.9	100.2	
286	Ghana	17.7	4.7	0.3	11.8	10.9	0.6	0.7	4.5	31	3.4	11.0	4.0	9.6	0.4	-	2.4	7.0	0.7	3.9	6.5	1.5	0.2	12.6	159	22.1	7.3	7.9	2.9	
372	Ireland	36.7	4.7	1.3	11.8	10.9	0.6	0.7	4.5	31	3.4	11.0	4.0	9.6	0.4	-	2.4	7.0	0.7	3.9	6.5	1.5	0.2	12.6	159	22.1	7.3	7.9	2.9	
380	Italy	162.7	25.2	15.8	224.1	152.9	27.1	74.3	37.1	691	61.9	82.7	86.6	50.8	13.1	5.7	46.8	91.9	52.2	30.1	64.2	142.1	35.3	169.1	3795	266.2	320.5	34.3	36.1	
392	Japan	1138.0	74.0	9.0	634.0	488.0	44.0	34.0	308.0	1680	257.0	581.0	179.0	222.0	21.0	12.0	152.0	456.0	69.0	70.0	36.0	38.0	123.0	889.0	1406.0	929.0	216.0	220.0	220.0	
400	Korea, Rep.	177.7	24.0	7.2	39.0	231.5	40.5	28.6	40.6	422	59.8	71.3	52.3	76.0	7.1	11.0	182.0	99.9	16.0	25.5	38.1	88.3	32.5	175.3	299.2	448.9	242.9	44.7	90.8	
410	Kuwait	2.5	0.9	-	0.2	8.4	-	-	0.4	3.7	1.7	2.8	4.9	0.5	-	4.4	17.0	26.9	7.5	24.4	31.0	47.4	18.1	51.6	33.2	9.9	11.7	0.5	1.3	
442	Mexico	97.0	89.0	5.2	68.8	29.7	-	23.5	4.9	72	33.1	11.2	57.3	66.2	-	-	4.4	17.0	26.9	7.5	24.4	31.0	47.4	18.1	51.6	33.2	9.9	11.4	2.6	6.4
486	Mongolia	12.0	0.4	-	7.3	11.8	3.3	4.1	6.8	22	24.4	60.7	57.6	31.9	0.8	-	5.6	25.6	6.3	6.2	36.8	3.4	0.2	0.8	786	112.2	57.1	0.2	1.2	
528	Netherlands	107.8	11.2	6.5	21.2	7.9	1.8	2.4	12.6	100	24.4	60.7	57.6	31.9	0.8	-	5.6	25.6	6.3	6.2	36.8	3.4	0.2	0.8	786	112.2	57.1	0.2	1.2	
578	New Zealand	45.9	3.4	0.6	15.4	12.2	0.5	0.4	15.2	68	11.8	32.8	8.3	5.4	1.2	0.2	1.2	5.7	0.3	1.5	4.2	7.4	10.4	21.2	330	16.3	23.3	1.9	2.5	
606	Poland	366.0	29.0	10.0	286.0	160.0	27.0	85.0	65.0	770	42.0	40.0	108.0	55.0	16.0	13.0	34.0	41.0	24.0	45.0	102.0	136.0	30.0	196.0	4160	236.0	276.0	37.0	62.0	
620	Portugal	99.8	18.0	1.8	158.3	145.3	9.8	59.4	54.9	381	18.5	33.6	12.4	23.8	-	-	6.6	16.8	25.6	10.6	37.8	15.1	5.2	79.7	40.5	41.1	39.5	4.2	15.0	
642	Romania	259.3	6.2	414.3	258.0	126.7	-	-	94.0	2036	43.3	26.3	182.8	-	-	-	86.3	175.9	-	-	-	173.2	0.8	189.1	6025	171.1	346.1	50.2	10.1	
702	Singapore	10.6	2.4	0.7	3.4	27.7	0.7	0.8	4.6	65	4.6	15.8	3.8	5.2	3.3	-	1.6	14.9	0.8	0.9	4.4	16.7	0.8	28.4	673	10.8	36.1	8.3	6.5	
702	Singapore	10.6	2.4	0.7	3.4	27.7	0.7	0.8	4.6	65	4.6	15.8	3.8	5.2	3.3	-	1.6	14.9	0.8	0.9	4.4	16.7	0.8	28.4	673	10.8	36.1	8.3	6.5	
718	South Africa	204.2	18.3	17.5	269.4	24.9	3.2	9.6	8.5	58	23.3	14.3	50.8	50.3	17.2	1.2	2.8	18.3	8.4	14.1	51.6	51.3	29.6	49.7	396	31.0	62.6	6.6	1.5	
826	United Kingdom	503.0	69.0	13.0	224.0	210.0	17.0	45.0	78.0	1140	148.0	304.0	144.0	170.0	9.0	8.0	65.0	169.0	44.0	42.0	123.0	151.0	59.0	337.0	5570	580.0	531.0	85.0	71.0	
840	United States	1333.0	137.0	41.0	829.0	807.0	48.0	67.0	508.0	4380	590.0	1538.0	402.0	485.0	72.0	40.0	205.0	670.0	38.0	141.0	364.0	413.0	247.0	1297.0	20660	151.0	894.0	902.0	369.0	

Source: UNIDO Database, FRED and author's calculations.

TABLE 3
WAGES AND SALARIES FOR 1990
(Millions of US dollars)

UNIDO Code	Country	Food Products (11/12)	Beverage (13)	Tobacco (14)	Textile (15)	Apparel (16)	Leather (17)	Footwear (18)	Wood (19)	Furniture (20)	Paper (21)	Printing (22)	Chemicals (23)	Other Chemicals (24)	Refined Petroleum (25)	Misc. Petroleum (26)	Rubber (27)	Plastic (28)	Metals (29)	Glass (30)	Other Metals (31)	Iron & Steel (32)	Non-ferrous Metals (33)	Ferrous Metals (34)	Machinery (35)	Electrical (36)	Transport Equipment (37)	Professional & Scientific Equipment (38)	Other Manufactures (39)					
40	Austria	12078	3660	481	748.7	346.2	48.9	132.1	387.5	576.0	616.6	760.8	643.0	54.8	2.6	1770	371	220.8	291.8	701	21.8	76.8	1072.5	266.4	1428.4	2317.7	8120	1357	1605	..				
52	Barbados	180	85	..	0.6	5.3	22.2	578.2	..	425.8	877.0	245.6	..	1247	28.3	118.1	421.4	30.9	502.6	1458.4			
56	Belgium	14531	3153	1147	869.2	433.2	35.0	22.2	578.2	..	425.8	877.0	245.6	..	1247	28.3	118.1	421.4	30.9	502.6	1458.4			
100	Bulgaria	1900	457	279	191.7	97.2	21.4	32.0	45.2	38.9	24.5	17.2	71.0	97.7	..	420	42.0	35.8	23.8	11.6	65.2	1458.4			
124	Canada	48330	8116	2074	1541.5	1541.5	93.4	200.6	2789.3	1330.7	3975.5	4012	1327.6	208.1	666.5	763.3	791.1	1399.6	300	31.7	1186.6	31.7	1186.6	1957.8	1023.3	3480.0	6691.4	489.5	973.6			
128	Chile	114			
196	Cyprus	30	238	64	16.1	70.1	67	16.5	16.5	19.4	7.5	20.3	10	11.9	3.8	..	2.2	11.3	11.6	07	07	30.6	20.2	3.9	24.5	12.9	6.4	44	00	8.7	..			
200	Czechoslovakia	3904	758	89	375.1	162.1	49.1	144.8	114.8	144.8	94.5	59.6	212.3	56.3	60.7	56.3	63.5	162	195	152	12.7	152	12.7	140.1	58.5	347.1	1144.3	389.9	458.8	44.6	90.3	..		
208	Denmark	22848	2542	62.4	404.1	222.5	20.2	42.5	365.3	802.4	353.6	434.9	367.1	634.6	278	992	83.1	330.3	545	76.4	49.7	132.4	132.4	243.6	33.3	132.4	983.3	779.3	502.9	271.1		
218	Ecuador	9.6	145	31	34.2	5.1	1.3	2.7	7.1	6.5	12.4	12.1	7.6	25.5	2.4	..	1.9	9.6	13.3	31	2.1	20.8	6.6	1.4	18.5	1.4	11.8	6.5	2.3	1.4		
242	Fiji	27.2	3.3	..	12.1	6.6	2.1	1.2	3.3		
246	Finland	12008	1752	388	234.5	279.8	28.4	63.1	748.1	253.0	153.1	118.9	449.2	304.3	..	1214	23.2	66.9	204.5	331	91.6	452.7	411.8	141.2	896.2	1627.3	811.6	854.1	181.6		
250	France	14005	19622	2411	4698.6	4192.8	1038.7	2579.1	2885.9	3800.1	8889.9	12774.4	7544.8	1887.3	..	1887.3	..	2062.9	7977.6	893.3	1887.3	469.2	819.1	234.9	1347.9	1611.8	1723.1	1886.5	258.6	2786.6		
252	Germany, West	27000	5413	4698.6	4192.8	4192.8	1038.7	2579.1	2885.9	3800.1	8889.9	12774.4	7544.8	1887.3	..	1887.3	..	2062.9	7977.6	893.3	1887.3	469.2	819.1	234.9	1347.9	1611.8	1723.1	1886.5	258.6	2786.6		
348	Hungary	4402	608	410	155.1	106.5	240	50.7	34.5	54.2	38.5	66.3	115.8	84.5	26.8	27.3	49.4	30.3	37.7	75.5	162.1	40.3	3.5	229.2	352.3	426.2	1930	162.8	46.1	
372	Ireland	7342	1561	410	192.2	126.5	88	10.3	74.5	41.2	79.5	276.6	148.0	244.3	13.2	55.4	133.3	116	92.5	121.3	40.3	3.5	229.2	352.3	426.2	1930	162.8	46.1		
380	Italy	59027	10867	42.4	687.5	381.6	752.0	1625.1	1091.1	1957.2	2275.3	3937.1	366.6	2236.0	704.4	2203	167.6	300.6	1853.8	109.7	214.3	5358.5	132.1	5489.5	1441.9	9982.2	11783.7	1893.6	1391.4		
392	Japan	213827	18579	585.6	1187.1	630.5	801.2	669.5	616.3	3625.9	6982.2	1846.2	688.7	7763.0	944.5	3789	427.4	10725.9	1887.7	214.0	8294.8	1309.6	390.5	2565.7	4759.9	46432.8	31328.1	5877.7	4993.4		
410	Korea, Rep.	13844	2614	106.5	2750.5	1496.8	344.7	201.8	345.2	349.1	580.0	83.2	719.9	812.4	148.8	117.8	1338.3	845.1	1157	30.4	834.9	1188.4	362.9	1655.5	2152.1	4101.8	3286.7	379.9	702		
414	Kuwait	32.9	84	..	4.8	16.2	5.1	15.8	4.2	14.0	34.7	4.6	..	1770	0.2	0.6	4.5	..	2.4	23.4	3.1	26.5	117.2	150	3.6	65		
442	Lebanon	46.8	256	..	26.2	89.7	..	84.7	7.2	24.1	49.5	54.6	164.6	47.9	128.6	105.4	85.7	186.7	109.7	388.2	97.2	183.3	392.9	98		
446	Malaysia	33.3	33.1	..	18.5	15.4	15.4	15.4	15.4	36.9	0.9	128.6	105.4	85.7	186.7	109.7	388.2	97.2	183.3	392.9	98	
496	Nepal	15.3	
538	Netherlands	29012	4256	208.7	321.7	146.1	41.2	48.3	296.7	224.6	702.5	1807.3	287.3	1015.4	308.1	593	158.7	678.2	1576	177.9	408.8	1073.6	1676.1	2065.5	3393.3	16289	1829	
554	New Zealand	11184	1382	244	133.2	41.5	10.7	8.1	386.1	163.3	351.5	899.2	303.4	171.7	53.4	29.6	33.7	153.8	192	42.7	192.5	239.3	375.3	585.2	1134.6	519.7	7189	562	660		
578	Norway	4610	407	151	315.2	157.5	28.4	84.1	78.1	82.3	52.7	46.4	157.4	75.3	40.2	23.7	38.5	49.1	270	53.8	120.4	237.8	54.2	229.7	544.0	288.0	3658	444	77.3	
616	Poland	4610	407	151	315.2	157.5	28.4	84.1	78.1	82.3	52.7	46.4	157.4	75.3	40.2	23.7	38.5	49.1	270	53.8	120.4	237.8	54.2	229.7	544.0	288.0	3658	444	77.3	
642	Romania	423.1	35.1	15.0	686.4	386.8	192.7	..	136.7	319.5	76.4	46.1	347.4	85.3	151.0	308.8	
662	Singapore	14.1	35.1	15.0	686.4	386.8	192.7	..	136.7	319.5	76.4	46.1	347.4	85.3	151.0	308.8	
702	Sri Lanka	38.9	316	184	75.7	27.1	22	19.5	15.5	30.9	24.4	39.8	23.1	33.4	13.0	13.4	10	4.4	20.1	14.7	56.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5	
722	Taiwan	34.9	376	456	457.1	28.5	63	15.1	13.1	9.6	36.6	37.7	146.9	133.4	84.8	31	5.1	29.1	145	27.7	107.5	147.1	56.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5	
788	Taiwan	34.9	376	456	457.1	28.5	63	15.1	13.1	9.6	36.6	37.7	146.9	133.4	84.8	31	5.1	29.1	145	27.7	107.5	147.1	56.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5	86.5
826	United Kingdom	91943	16869	4173	3599.2	2365.2	269.9	671.2	1480.5	2319.0	3478.5	8288.8	4344.2	4456.9	330.3	220.2	1413.4	3546.0	7582	988.8	29263	36255	36255	36255	36255	36255	36255	36255	36255	36255	36255	36255	36255	36255
840	United States	292300	4240	15200	15130.0	10820.0	860.0	990.0	10090.0	8160.0	18800.0	38800.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0	15200.0

Source: UNIDO Database, FRED and author's calculations.

TABLE 4
VALUE ADDED FOR 1990
(Millions of US Dollars)

UNIDO Code	Country	Food Products (311/312)	Beverage (313)	Tobacco (314)	Textile (321)	Apparel (322)	Leather (323)	Footwear (324)	Wood (331)	Furniture (332)	Paper Products (341)	Printing and Publishing (342)	Chemicals (351)	Other Chemicals (352)	Petroleum Refineries (353)	Misc. Products of Petroleum & Coal (354)	Rubber (355)	Plastic (356)	Pottery (361)	Glass (362)	Other Non- Metallic Mineral Products (369)	Iron & Steel (371)	Non- Ferrous Metals (372)	Fabricated Metal Products (381)	Machinery except Electrical (382)	Electrical Machinery (383)	Transport Equipment (384)	Professional & Scientific Equipment (385)	Other Manufacturing Industries (390)	
40	Austria	2302.3	840.6	1417.4	1291.1	547.1	82.1	213.1	879.1	993.9	1333.3	1162.9	1277.0	1069.6	488.8	64.6	311.0	544.6	111.6	517.6	1472.8	2088.4	434.4	2533.6	3292.2	3926.5	1651.9	221.6	248.9	
52	Barbados	29.9	11.2	0.6	5.4	2.5	..	10.8	..	9.3	4.1	10.3	3.9	..	0.8	..	0.6	
56	Belgium	6044.2	846.0	310.0	2131.5	917.8	55.6	..	321.3	1613.3	1042.2	1677.0	4482.9	1198.5	473.4	..	272.0	..	1782.4	2197.0	1140.2	13530.2	2458.7
100	Bulgaria	27.5	2.2	1.5	3.6	2.2	2.9	0.4	0.7	1.3	0.2	3.2	1.0	..	0.1	1.1	..	0.4	
124	Canada	12701.7	2948.3	977.1	2974.0	2828.3	162.8	334.3	4465.3	2245.5	8750.7	7670.7	4808.1	6236.6	2271.2	291.4	1397.0	2896.9	68.6	642.8	2802.6	3231.1	3222.6	6453.7	7576.5	7465.1	14124.5	925.6	1705.6	
170	Colombia	1306.1	927.8	173.2	816.3	221.0	65.7	99.6	54.3	37.8	300.6	212.0	521.6	597.3	151.3	33.8	131.4	223.0	59.7	113.5	338.5	280.7	55.7	278.7	123.6	270.8	332.5	69.9	84.1	
196	Cyprus	100.8	73.0	40.9	32.1	117.6	10.8	29.7	39.1	35.8	16.8	37.0	3.2	27.4	6.8	..	3.0	25.1	1.7	1.5	69.1	54.8	23.8	11.8	8.8	0.1	19.3	
200	Czechoslovakia	916.4	257.9	24.0	790.0	223.4	65.7	255.7	288.6	154.3	254.6	127.0	698.1	176.6	315.9	209.5	131.5	49.0	45.7	298.1	411.1	1270.8	236.2	601.7	2596.7	894.2	903.1	83.6	192.2	
208	Denmark	4072.3	757.0	203.1	610.3	259.0	24.7	65.0	485.6	642.0	628.1	191.8	1106.7	1537.2	117.6	207.0	122.0	635.4	70.8	113.9	941.1	281.0	72.6	1837.4	3050.4	1318.9	1127.9	622.3	489.1	
218	Ecuador	228.1	33.4	1.4	94.5	10.4	3.6	5.5	15.9	9.4	34.5	26.8	17.5	75.3	374.0	3.9	17.1	42.0	7.0	8.1	59.7	18.6	2.5	43.9	2.8	32.4	21.8	2.9	3.5	
242	Fiji	60.2	13.9	..	16.0	..	0.8	..	10.8	3.1	5.1	6.2	..	7.3	0.9	3.5	5.0	5.5	1.2	..	1.4	..	1.2	
246	Finland	2575.9	665.9	176.5	386.0	428.4	47.6	92.6	1578.7	515.0	3603.5	2113.8	1371.5	706.9	674.5	121.1	133.1	425.3	73.5	162.7	1053.2	850.3	363.0	1759.6	3355.0	1832.6	1405.5	344.2	168.2	
250	France	25554.1	5382.1	1919.1	7666.1	5806.8	1129.6	1419.8	4182.3	3972.8	6822.6	12499.8	10872.1	12426.7	15128.6	..	3340.9	6662.4	3089.6	7522.8	8433.5	4533.3	20995.5	24819.2	25769.6	28616.1	41099.0	4318.6		
280	Germany, West	28389.1	11911.0	12633.2	11848.7	5886.7	944.4	1152.5	6178.9	7884.9	13489.8	10255.2	35536.8	27941.2	19675.5	..	6413.6	17312.5	1554.6	4790.8	12031.0	19204.2	7733.3	39180.5	82542.1	72566.2	67433.2	80103.8	2849.2	
348	Hungary	739.6	165.7	42.4	299.4	220.1	37.1	84.4	88.1	119.9	128.5	189.7	404.3	436.6	460.4	..	79.2	144.8	54.8	82.2	199.4	337.5	201.0	275.3	756.6	666.8	396.1	292.6	101.2	
372	Ireland	3059.6	789.5	165.1	347.7	206.3	21.0	18.5	169.9	85.8	189.9	59.2	755.4	1713.4	30.3	..	117.8	330.6	28.0	144.1	558.1	91.5	9.9	468.1	2229.4	1835.0	308.1	609.7	131.7	
380	Italy	9598.5	2014.9	535.9	10272.2	4876.1	1234.5	2231.0	1615.9	2900.4	3877.8	6171.4	5906.0	3973.8	1717.7	406.5	2254.4	4799.3	2859.5	1673.5	4299.3	8117.0	1787.8	8013.5	20329.7	14990.4	14549.7	1761.1	1890.5	
392	Japan	66675.9	10046.2	2002.9	27046.1	11920.7	1864.8	1478.0	14006.5	8729.9	22287.5	47938.4	38075.8	46760.3	4841.5	1540.2	11402.7	30796.3	2983.6	8467.4	26652.4	48539.3	11976.0	62904.9	126562.6	133883.6	95599.6	12797.8	13730.2	
410	Korea, Rep	6046.7	1889.1	2794.0	6832.8	3400.7	1144.5	593.8	875.6	972.2	2122.6	2530.5	4181.5	2925.7	2865.1	517.0	3062.6	2734.0	274.5	991.3	3697.3	6187.3	1200.8	5144.5	7004.4	15065.8	10242.2	1143.9	1769.0	
414	Kuwait	70.0	21.0	..	15.8	54.7	10.1	30.3	31.0	5.5	43.9	15.0	1672.0	0.4	1.9	16.3	..	12.4	73.2	10.7	..	55.1	18.8	27.8	2.3	..	16.8	
442	Luxembourg	66.8	73.5	..	96.2	15.2	..	86.2	..	354.3	28.0	235.4	..	623.0	61.9	209.6	157.6	157.6	103.1	18.8	
484	Mexico	2177.5	2209.6	679.8	765.7	217.0	..	142.4	53.4	68.0	705.7	170.0	2089.8	2025.1	..	172.3	464.1	441.1	145.5	626.9	985.8	1799.0	706.4	945.2	745.5	1540.5	3384.7	81.3	96.2	
496	Mongolia	71.1	51.1	..	65.6	30.4	21.9	20.4	21.6	0.4	..	7.9	..	5.1	0.8	..	22.2	0.6	..	16.2	..	3.1	..	0.7	29.2	
528	Netherlands	6036.4	1500.3	1847.9	1002.2	233.9	69.7	72.5	477.2	361.9	1618.4	3217.0	5591.6	1845.7	1095.0	131.3	284.5	1304.8	303.7	358.1	1016.0	2108.3	..	2903.4	3551.5	5286.3	2464.1	308.1	110.9	
554	New Zealand	1675.2	215.4	44.7	232.1	202.2	54.3	40.6	322.8	125.9	552.4	536.9	248.8	211.2	137.2	8.9	62.0	228.5	256.5	112.8	139.0	479.7	339.5	259.5	32.2	23.9
578	Norway	1307.1	660.3	478.1	191.2	58.1	15.8	11.2	619.0	236.1	787.3	1380.7	811.4	392.8	194.7	63.3	57.5	278.1	27.0	77.3	360.6	346.7	826.1	783.9	1590.3	750.5	1027.7	81.8	89.3	
616	Poland	2595.2	1838.4	379.3	1221.9	432.2	120.1	262.5	324.7	306.9	348.3	165.7	1055.9	649.4	1418.9	248.7	209.4	273.8	107.5	226.8	602.4	1886.7	950.6	1080.6	2604.3	1420.2	1855.2	172.7	258.3	
620	Portugal	1305.2	343.8	592.0	1654.2	984.6	126.0	452.1	532.1	232.6	576.7	522.8	432.3	480.9	54.2	236.5	291.3	173.4	723.6	272.5	80.7	825.5	528.4	833.6	582.9	35.5	141.2	
642	Romania	1649.4	641.9	410.1	1448.8	691.0	66.9	365.5	312.1	321.0	169.4	142.7	111.4	441.3	133.7	35.7	129.3	396.8	570.6	120.4	40.1	650.9	22.3	869.3	2434.0	1208.1	704.4	..	227.4	
702	Singapore	321.9	138.9	64.0	71.6	294.0	10.9	9.1	54.5	89.3	189.2	54.5	383.8	600.3	97.1	..	24.7	328.7	33.5	..	148.6	96.9	40.6	148.6	2737.3	2707.6	890.3	200.5	1144.4	
716	Zimbabwe	237.1	301.9	75.9	255.0	102.1	7.2	66.5	43.3	31.6	64.3	93.5	114.9	126.9	36.6	47.0	2.6	9.1	54.4	184.2	12.7	135.4	43.2	87.7	80.7	2.5	12.7	
818	Egypt	949.2	46.1	133.5	847.1	52.9	6.7	22.2	16.8	13.0	84.0	76.0	338.5	480.4	1155.9	35.9	7.6	72.5	62.1	41.8	330.5	313.7	175.6	144.5	142.1	160.5	237.4	37.3	82	
836	United Kingdom	25001.0	6605.4	2361.6	6996.0	4652.2	532.7	1260.7	3196.2	4527.9	7990.4	19352.0	14698.6	14888.8	4403.6	745.8	3000.8	8203.5	14560.0	2077.5	8984.7	8043.7	2770.0	14933.1	29901.8	22231.0	28763.1	3640.1	2770.0	
840	United States	119830.0	21140.0	22560.0	54960.0	25480.0	2210.0	2320.0	20850.0	16910.0	57200.0	103180.0	73480.0	81770.0	22820.0	4890.0	13430.0	37320.0	1840.0	10080.0	23980.0	31780.0	17510.0	73360.0	145060.0	112400.0	154030.0	76520.0	18720.0	

Source: UNIDO Database, FRED and author's calculations.

TABLE 5
ELASTICITY OF SUBSTITUTION BETWEEN LABOR AND CAPITAL
CES Production Function

Industry (3 digit ISIC)	($\rho - 1$)	se	n	R ²	σ	p-value ($\rho=0$)	σ^*	se*
Food Products (311/312)	-1.324	0.21	31	0.58	0.76	0.14	0.83	0.18
Beverage (313)	-1.161	0.26	24	0.47	0.86	0.54	0.95	0.18
Tobacco (314)	-0.471	0.41	18	0.08	2.12	0.21	1.68	0.28
Textile (321)	-1.076	0.18	28	0.58	0.93	0.67	0.96	0.19
Apparel (322)	-1.421	0.23	25	0.63	0.70	0.08	1.18	0.16
Leather (323)	-1.159	0.24	20	0.57	0.86	0.51	1.53	0.19
Footwear (324)	-1.733	0.17	21	0.84	0.58	0.00	0.89	0.26
Wood (331)	-1.414	0.16	29	0.74	0.71	0.02	0.92	0.16
Furniture (332)	-1.233	0.15	22	0.77	0.81	0.14	1.03	0.18
Paper Products (341)	-1.243	0.16	29	0.69	0.80	0.14	0.89	0.20
Printing and Publishing (342)	-1.472	0.21	26	0.67	0.68	0.04	1.07	0.25
Chemicals (351)	-1.255	0.18	26	0.67	0.80	0.17	1.24	0.23
Other Chemicals (352)	-1.346	0.21	22	0.68	0.74	0.11	0.74	0.18
Petroleum Refineries (353)	-0.928	0.27	18	0.43	1.08	0.79	1.34	0.18
Misc. Products of Petroleum & Coal (354)	-1.083	0.31	12	0.56	0.92	0.79	1.79	0.38
Rubber (355)	-1.522	0.17	24	0.78	0.66	0.01	0.90	0.24
Plastic (356)	-1.691	0.19	20	0.81	0.59	0.00	0.67	0.37
Pottery (361)	-1.207	0.15	22	0.76	0.83	0.19	2.82	0.17
Glass (362)	-1.044	0.17	18	0.69	0.96	0.81	1.12	0.22
Other Non-Metallic Mineral Products (369)	-1.582	0.19	21	0.78	0.63	0.01	0.71	0.33
Iron & Steel (371)	-1.066	0.32	25	0.32	0.94	0.84	1.23	0.29
Non-Ferrous Metals (372)	-1.514	0.26	18	0.68	0.66	0.06	0.87	0.19
Fabricated Metal Products (381)	-1.094	0.13	30	0.71	0.91	0.49	1.34	0.18
Machinery except Electrical (382)	-1.046	0.27	22	0.43	0.96	0.87	0.82	0.21
Electrical Machinery (383)	-1.452	0.20	24	0.71	0.69	0.03	0.78	0.20
Transport Equipment (384)	-1.131	0.23	23	0.54	0.88	0.57	0.93	0.22
Professional & Scientific Equipment (385)	-0.982	0.21	20	0.55	1.02	0.93	3.51	0.26
Other Manufacturing Industries (390)	-0.726	0.20	23	0.39	1.38	0.18	1.85	0.20

Note:

se: standard error

n: Number of countries

p-value: probability that $\sigma = 1$.

* Estimate of σ ($=1/(1-\rho)$) and standard error of coefficient ($1-\rho$) using average w/r rather than sector-specific w/r.

FIGURE 1
CROSS-COUNTRY CORRELATIONS OF RELATIVE FACTOR PRICES AND
FACTOR INTENSITIES

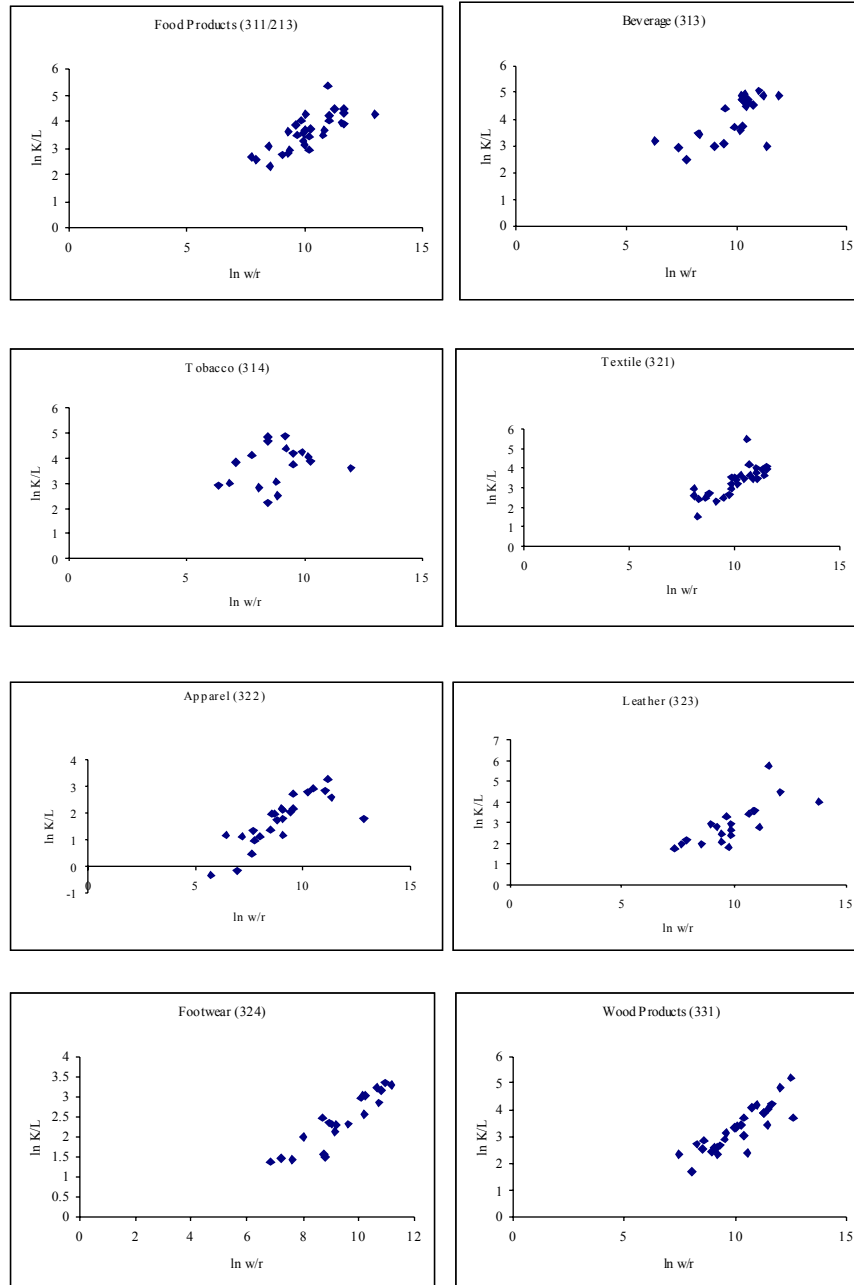


FIGURE 1
CROSS-COUNTRY CORRELATIONS OF RELATIVE FACTOR PRICES AND
FACTOR INTENSITIES

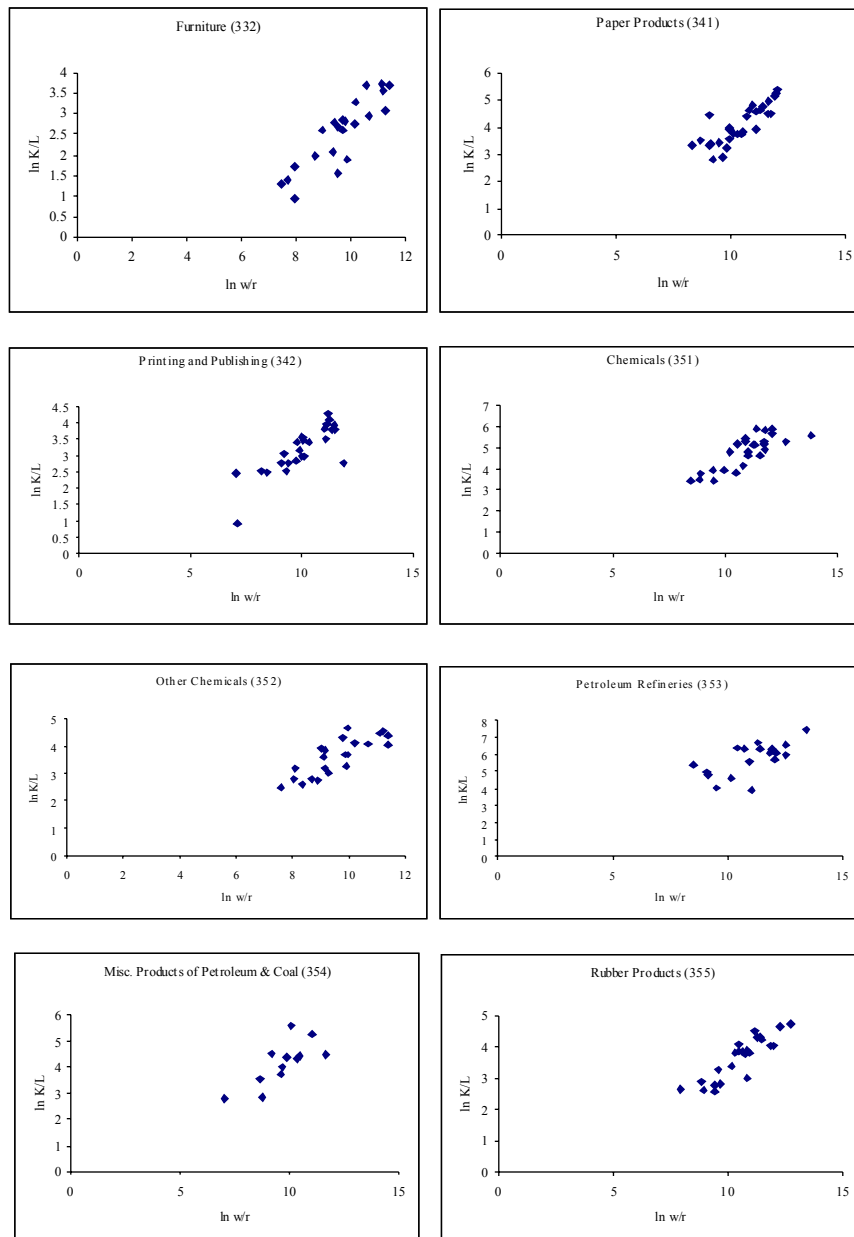


FIGURE 1
CROSS-COUNTRY CORRELATIONS OF RELATIVE FACTOR PRICES AND
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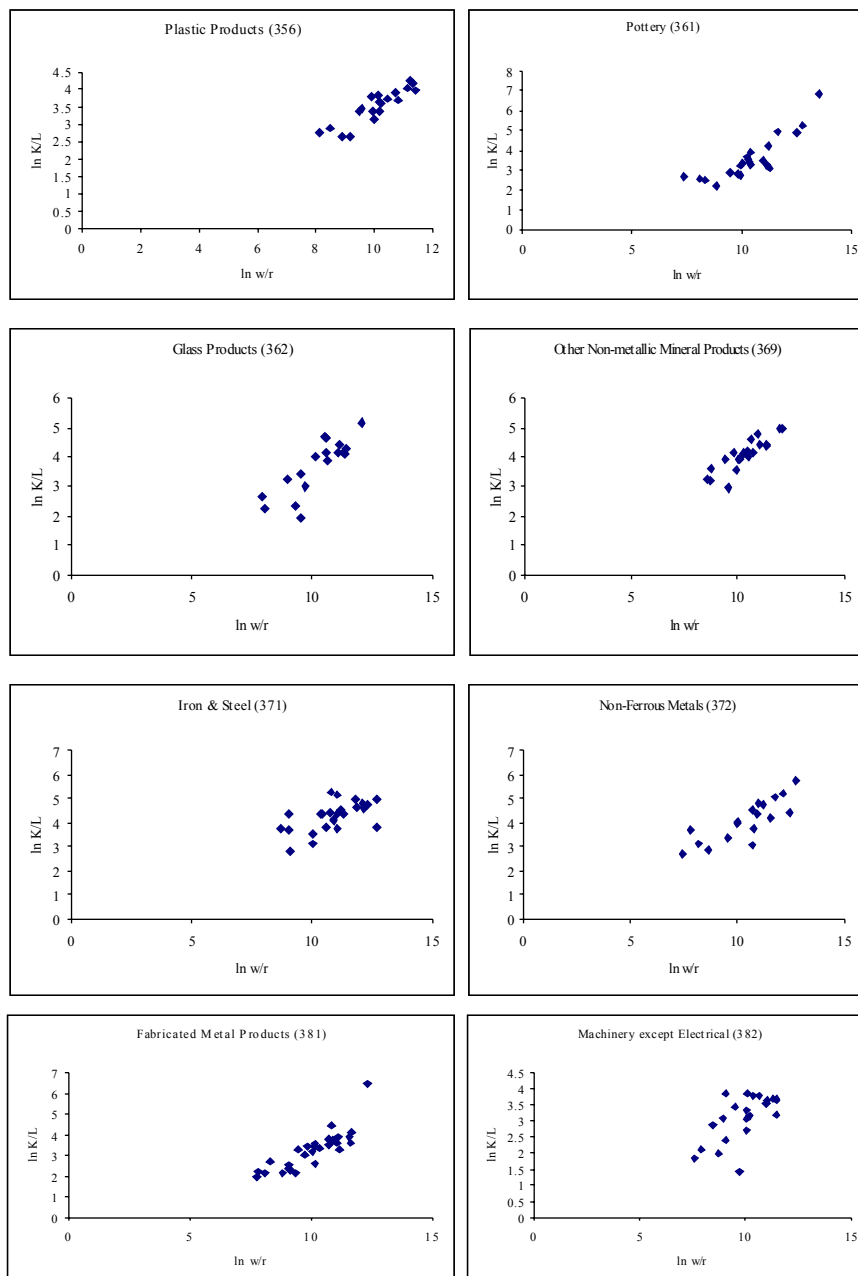


FIGURE 1
CROSS-COUNTRY CORRELATIONS OF RELATIVE FACTOR PRICES AND
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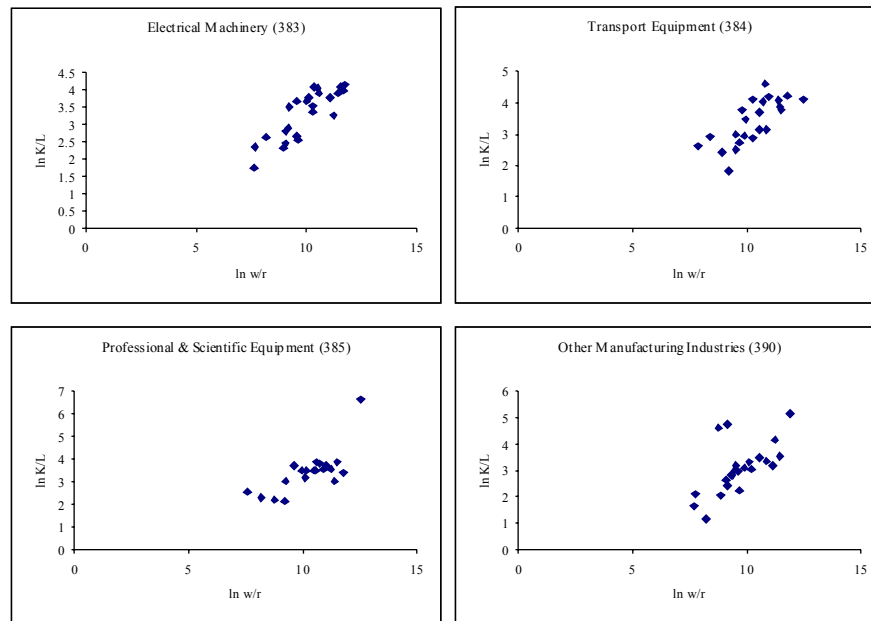


Table 6 presents a comparison of the estimates of σ reported in table 5 with those obtained by Behrman (1982). The first column in Table 6 replicates column 5 in Table 5. Column 2 presents the estimates of Behrman using pooled data for 70 countries. These coefficients come from an estimation of equation (6) that includes interactions of $\ln w_{ic}$ and sectorial dummies, to allow for cross-industry variations in σ . Its estimates are very close to the Cobb-Douglas value of 1.0. This procedure, however, does not include non-interactive dummies. This is therefore difficult to interpret the results as the constant term in (6) has a sector/country-specific component. Column 3 reports estimations of equation (6) using a similar procedure to Behrman's. The results are very similar to those in column 2, and they are subject to similar criticisms. Indeed, the correlation coefficient is .87. Although the coefficients are also close to 1, the null hypothesis of Cobb-Douglas technologies is rejected in almost all industries. The correlation coefficient between the coefficients in columns 1 and 3 is .57.

TABLE 6
COMPARISON OF ALTERNATIVE APPROACHES

Industry (3 digit ISIC)	Claro Eq. (4)	Behrman Eq. (6)	Claro Eq. (6)	Claro Eq. (6)*	Claro Eq. (6)**
Food Products (311/312)	0.76	0.91	0.93	0.93	0.84
Beverage (313)	0.86	0.97	0.98	0.91	0.89
Tobacco (314)	2.12	1.00	1.04	1.18	0.95
Textile (321)	0.93	0.88	0.91	0.98	0.82
Apparel (322)	0.70	0.87	0.90	0.83	0.79
Leather (323)	0.86	0.87	0.91	0.94	0.80
Footwear (324)	0.58	0.86	0.91	0.80	0.79
Wood (331)	0.71	0.88	0.90	0.92	0.81
Furniture (332)	0.81	0.86	0.89	1.11	0.80
Paper Products (341)	0.80	0.91	0.94	1.01	0.85
Printing and Publishing (342)	0.68	0.87	0.91	0.92	0.82
Chemicals (351)	0.80	0.94	0.95	1.24	0.86
Other Chemicals (352)	0.74	0.94	0.97	0.78	0.87
Petroleum Refineries (353)	1.08	0.99	1.04	0.73	0.94
Misc. Products of Petroleum & Coal (354)	0.92	0.89	0.98	0.66	0.87
Rubber (355)	0.66	0.91	0.90	1.16	0.81
Plastic (356)	0.59	0.90	0.94	0.74	0.83
Pottery (361)	0.83	0.87	0.93	0.85	0.82
Glass (362)	0.96	0.88	0.93	0.84	0.83
Other Non-Metallic Mineral Products (369)	0.63	0.91	0.94	1.03	0.84
Iron & Steel (371)	0.94	0.89	0.93	0.75	0.83
Non-Ferrous Metals (372)	0.66	0.89	0.95	0.63	0.84
Fabricated Metal Products (381)	0.91	0.89	0.91	1.04	0.82
Machinery except Electrical (382)	0.96	0.86	0.92	0.90	0.81
Electrical Machinery (383)	0.69	0.89	0.92	1.07	0.82
Transport Equipment (384)	0.88	0.87	0.90	1.14	0.81
Professional & Scientific Equipment (385)	1.02	0.86	0.93	0.78	0.82
Other Manufacturing Industries (390)	1.38		0.93	0.85	0.82

* Includes non-interactive industry dummies.

** Includes non-interactive country dummies.

Columns 4 and 5 in table 6 report results of estimations of equation (6) that incorporate sector-specific or country-specific non-interactive dummies. The results

with sector-specific dummies are orthogonal to those obtained under other specifications. The results with country-specific non-interactive dummies are highly correlated with those in columns 2 and 3 and to less extent to those in column 1. Again, the Cobb-Douglas hypothesis is rejected in almost all industries. The rejection of the Cobb-Douglas hypothesis contrasts with other studies that have found evidence of $\sigma = 1$ (e.g., Lovell (1973) and Corbo and Meller (1982)). However, other studies summarized in Hamermesh (1993) find mix evidence on the elasticity of substitution, specially using aggregate data.

5. CONCLUSION

Many alternative methodologies have been proposed to estimate elasticities of substitution between labor and capital, σ_{LK} . However, most of them suffer from the ability to account for international technology differences in a wise manner. This paper proposes a simple methodology that is consistent with international hicks-neutral technology gaps and differences in returns to scale. It relies on sectorial data on capital stock, that is estimated for 34 developed and developing countries in 1990. The results show that most estimates of σ_{LK} are close to one, as most of the literature have founded. However, the null-hypothesis of Cobb-Douglas production functions is in general rejected in favor of production functions with $\sigma_{LK} < 1$.

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